## Lecture 9: Linear Regression

## Goals

- Develop basic concepts of linear regression from a probabilistic framework
- Estimating parameters and hypothesis testing with linear models
- Linear regression in R


## Regression

- Technique used for the modeling and analysis of numerical data
- Exploits the relationship between two or more variables so that we can gain information about one of them through knowing values of the other
- Regression can be used for prediction, estimation, hypothesis testing, and modeling causal relationships


## Regression Lingo



Dependent Variable Independent Variable
Outcome Variable Predictor Variable
Response Variable Explanatory Variable

## Why Linear Regression?

- Suppose we want to model the dependent variable $Y$ in terms of three predictors, $X_{1}, X_{2}, X_{3}$

$$
Y=f\left(X_{1}, X_{2}, X_{3}\right)
$$

- Typically will not have enough data to try and directly estimate f
- Therefore, we usually have to assume that it has some restricted form, such as linear

$$
Y=X_{1}+X_{2}+X_{3}
$$

## Linear Regression is a Probabilistic Model

- Much of mathematics is devoted to studying variables that are deterministically related to one another

$$
y=\beta_{0}+\beta_{1} x
$$



- But we're interested in understanding the relationship between variables related in a nondeterministic fashion


## A Linear Probabilistic Model

- Definition: There exists parameters $\beta_{0}, \beta_{1}$, and $\sigma^{2}$, such that for any fixed value of the independent variable $x$, the dependent variable is related to $x$ through the model equation

$$
y=\beta_{0}+\beta_{1} x+\varepsilon
$$

- $\varepsilon$ is a rv assumed to be $\mathrm{N}\left(0, \sigma^{2}\right)$



## Implications

- The expected value of $Y$ is a linear function of $X$, but for fixed $x$, the variable $Y$ differs from its expected value by a random amount
- Formally, let $x^{*}$ denote a particular value of the independent variable $\times$, then our linear probabilistic model says:

$$
\begin{aligned}
& E\left(Y \mid x^{*}\right)=\mu_{\mathrm{Ylx}^{*}}=\text { mean value of } Y \text { when } x \text { is } x * \\
& V\left(Y \mid x^{*}\right)=\sigma_{\mathrm{Y} \mid \mathrm{x}^{*}}^{2}=\text { variance of } Y \text { when } x \text { is } x *
\end{aligned}
$$

## Graphical Interpretation



- For example, if $x=$ height and $y=$ weight then $\mu_{Y \mid x=60}$ is the average weight for all individuals 60 inches tall in the population


## One More Example

Suppose the relationship between the independent variable height $(\mathrm{x})$ and dependent variable weight ( y ) is described by a simple linear regression model with true regression line

$$
y=7.5+0.5 x \text { and } \sigma=3
$$

- Q1: What is the interpretation of $\beta_{1}=0.5$ ?

The expected change in height associated with a 1 -unit increase in weight

- Q2: If $x=20$ what is the expected value of $Y$ ?

$$
\mu_{Y \mid \mathrm{x}=20}=7.5+0.5(20)=17.5
$$

- Q3: If $x=20$ what is $P(Y>22)$ ?

$$
P(\mathrm{Y}>22 \mid \mathrm{x}=20)=\mathrm{P}\left(\frac{22-17.5}{3}\right)=1-\phi(1.5)=0.067
$$

## Estimating Model Parameters

- Point estimates of $\hat{\beta}_{0}$ and $\hat{\beta}_{1}$ are obtained by the principle of least squares

$$
f\left(\beta_{0}, \beta_{1}\right)=\sum_{i=1}^{n}\left[y_{i}-\left(\beta_{0}+\beta_{1} x_{i}\right)\right]^{2}
$$



- $\hat{\beta}_{0}=\bar{y}-\hat{\beta}_{1} \bar{x}$


## Predicted and Residual Values

- Predicted, or fitted, values are values of $y$ predicted by the leastsquares regression line obtained by plugging in $\mathrm{x}_{1}, \mathrm{x}_{2}, \ldots, \mathrm{x}_{\mathrm{n}}$ into the estimated regression line

$$
\begin{aligned}
& \hat{y}_{1}=\hat{\beta}_{0}-\hat{\beta}_{1} x_{1} \\
& \hat{y}_{2}=\hat{\beta}_{0}-\hat{\beta}_{1} x_{2}
\end{aligned}
$$

- Residuals are the deviations of observed and predicted values

$$
\begin{aligned}
& e_{1}=y_{1}-\hat{y}_{1} \\
& e_{2}=y_{2}-\hat{y}_{2}
\end{aligned}
$$



## Residuals Are Useful!

- They allow us to calculate the error sum of squares (SSE):

$$
S S E=\sum_{i=1}^{n}\left(e_{i}\right)^{2}=\sum_{i=1}^{n}\left(y_{i}-\hat{y}_{i}\right)^{2}
$$

- Which in turn allows us to estimate $\sigma^{2}$ :

$$
\hat{\sigma}^{2}=\frac{S S E}{n-2}
$$

- As well as an important statistic referred to as the coefficient of determination:

$$
r^{2}=1-\frac{S S E}{S S T} \quad S S T=\sum_{i=1}^{n}\left(y_{i}-\bar{y}\right)^{2}
$$

## Multiple Linear Regression

- Extension of the simple linear regression model to two or more independent variables

$$
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\ldots+\beta_{n} x_{n}+\varepsilon
$$

$$
\text { Expression }=\text { Baseline }+ \text { Age }+ \text { Tissue }+ \text { Sex }+ \text { Error }
$$

- Partial Regression Coefficients: $\beta_{\mathrm{i}} \equiv$ effect on the dependent variable when increasing the $i^{\text {th }}$ independent variable by 1 unit, holding all other predictors constant


## Categorical Independent Variables

- Qualitative variables are easily incorporated in regression framework through dummy variables
- Simple example: sex can be coded as 0/1
- What if my categorical variable contains three levels:



## Categorical Independent Variables

- Previous coding would result in colinearity
- Solution is to set up a series of dummy variable. In general for $k$ levels you need $k-1$ dummy variables

$$
\begin{aligned}
& x_{1}=\left\{\begin{array}{l}
1 \text { if AA } \\
0 \text { otherwise }
\end{array}\right. \\
& x_{2}=\left\{\begin{array}{l}
1 \text { if AG } \\
0 \text { otherwise }
\end{array}\right.
\end{aligned}
$$

|  | $x_{1}$ | $x_{2}$ |
| :---: | :---: | :---: |
| AA | 1 | 0 |
| AG | 0 | 1 |
| GG | 0 | 0 |

## Hypothesis Testing: Model Utility Test (or Omnibus Test)

- The first thing we want to know after fitting a model is whether any of the independent variables ( $X$ 's) are significantly related to the dependent variable $(\mathrm{Y})$ :

$$
\begin{aligned}
& \mathrm{H}_{0}: \beta_{1}=\beta_{2}=\ldots=\beta_{k}=0 \\
& \mathrm{H}_{\mathrm{A}}: \text { At least one } \beta_{1} \neq 0 \\
& f=\frac{R^{2}}{\left(1-R^{2}\right)} \cdot \frac{k}{n-(k+1)}
\end{aligned}
$$

Rejection Region: $F_{\alpha, k, n-(k+1)}$

## Equivalent ANOVA Formulation of Omnibus Test

- We can also frame this in our now familiar ANOVA framework
- partition total variation into two components: SSE (unexplained variation) and SSR (variation explained by linear model)


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| Source of <br> Variation | df | Sum of Squares | MS | F |
| :---: | :---: | :---: | :---: | :---: |
| Regression | k | $S S R=\sum\left(\hat{y}_{i}-\bar{y}\right)^{2}$ | $\frac{S S R}{k}$ | $\frac{M S_{R}}{M S_{E}}$ |
| Error | $\mathrm{n}-2$ | $S S E=\sum\left(y_{i}-\hat{y}_{i}\right)^{2}$ | $\frac{S S E}{n-2}$ |  |
| Total | $\mathrm{n}-1$ | $S S T=\sum\left(y_{i}-\bar{y}\right)^{2}$ |  |  |

Rejection Region : $F_{\alpha, k, n-(k+1)}$

## F Test For Subsets of Independent Variables

- A powerful tool in multiple regression analyses is the ability to compare two models
- For instance say we want to compare:

$$
\begin{aligned}
\text { Full Model: } y & =\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{4} x_{4}+\varepsilon \\
\text { Reduced Model: } y & =\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\varepsilon
\end{aligned}
$$

- Again, another example of ANOVA:
$S S E_{R}=$ error sum of squares for reduced model with $l$ predictors
$\mathrm{SSE}_{\mathrm{F}}=$ error sum of squares for

$$
f=\frac{\left(S S E_{R}-S S E_{F}\right) /(k-l)}{S S E_{F} /([n-(k+1)]}
$$

full model with $k$ predictors

## Example of Model Comparison

- We have a quantitative trait and want to test the effects at two markers, M1 and M2.

Full Model: Trait $=$ Mean $+M 1+M 2+(M 1 * M 2)+$ error Reduced Model: Trait $=$ Mean + M1 + M2 + error

$$
f=\frac{\left(S S E_{R}-S S E_{F}\right) /(3-2)}{S S E_{F} /([100-(3+1)]}=\frac{\left(S S E_{R}-S S E_{F}\right)}{S S E_{F} / 96}
$$

Rejection Region: $F_{a, 1,96}$

## Hypothesis Tests of Individual Regression Coefficients

- Hypothesis tests for each $\hat{\beta}_{i}$ can be done by simple t-tests:

$$
\begin{gathered}
\qquad \begin{array}{c}
\mathrm{H}_{0}: \hat{\beta}_{i}=0 \\
\mathrm{H}_{\mathrm{A}}: \hat{\beta}_{i} \neq 0 \\
\mathrm{~T}=\frac{\hat{\beta}_{i}-\beta_{i}}{\operatorname{se}\left(\beta_{i}\right)}
\end{array} \\
\text { Critical value }: t_{\alpha / 2, n-(k-1)}
\end{gathered}
$$

- Confidence Intervals are equally easy to obtain:

$$
\hat{\beta}_{i} \pm t_{\alpha / 2, n-(k-1)} \bullet \operatorname{se}\left(\hat{\beta}_{i}\right)
$$

## Checking Assumptions

- Critically important to examine data and check assumptions underlying the regression model
$>$ Outliers
$>$ Normality
$>$ Constant variance
$>$ Independence among residuals
- Standard diagnostic plots include:
$>$ scatter plots of $y$ versus $x_{i}$ (outliers)
$>$ qq plot of residuals (normality)
$>$ residuals versus fitted values (independence, constant variance)
$>$ residuals versus $\mathrm{x}_{\mathrm{i}}$ (outliers, constant variance)
- We'll explore diagnostic plots in more detail in R


## Fixed -vs- Random Effects Models

- In ANOVA and Regression analyses our independent variables can be treated as Fixed or Random
- Fixed Effects: variables whose levels are either sampled exhaustively or are the only ones considered relevant to the experimenter
- Random Effects: variables whose levels are randomly sampled from a large population of levels
- Example from our recent AJHG paper:
Expression = Baseline + Population + Individual + Error

